### Credit and House Prices Cycles

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# Chapter 1: Credit and House Prices Cycles

### Introduction

#### Motivation

- The study of housing prices and excessive credit has become more important in understanding financial market stability
- We also observed increasing use of monetary policies, significant growth in macro balance sheet size, including real estate values and total credit lending to household
- We study the dynamic relationship between housing prices and household credit in this paper

## Contribution

- 1. Relationship between housing prices and household credit
- Apply Unobserved Component Model (Clark 1987) to extract information about trends and cycles

 $\Rightarrow$  Jointly examine the two variables and their interaction both in the long-run and short-run

• Specify cycles to be VAR process (cross-cycle) rather than univariate AR process

 $\Rightarrow$  Test if past movement of one cycle has predictive power over another cycle

# Contribution

- 2. Technical contribution to the optimization process:
- Novel numerical optimization / parameters constraint method to ensure the cyclical components are in feasible stationary region
- 3. Overcome "curse of dimensionality" using Bayesian method:
  - Common problem in estimating complex unobserved component state space model
  - We use random walk Metropolis-Hasting method to estimate posterior distribution of parameters of interest

#### Literature Review

- 1. Dynamics of credit changes:
- Kiyotaki & Moore (1997), Myerson (2012), Guerrieri & Uhlig (2016), Boissay et al (2016).
- 2. Dynamics of house prices changes:
- Hong & Stein (1999), Glaeser et al (2008) (2017), Kishor, Kumari, & Song (2015)

#### Literature Review

- 3. House price cycles generates credit cycles:
- Bernanke & Gertler (1989), Bernanke et al (1999); Kiyotaki & Moore (1997) "
- Empirical Evidence: Fitzpatrick and McQuinn (2007), Berlinghieri (2010), Gimeno and Martinez-Carrascal (2010), Anundsen and Jansen (2013), for evidence from Ireland, USA, Spain and Norway, respectively
- 4. Credit cycles genereates house price cycles:
  - Agnello & Schuknecht (2011), Kermani (2012), Justiniano et al (2019), Schularick et al (2012) (2016)

 $\Rightarrow$  However, the debate on which cycle causes changes on the other is still open

#### Data

Bank of International Settlement (BIS)

- Household Credit to GDP: Total Credit to non-financial sector (household)
- House Price Index: Residential property prices: selected series (real value). Index = 100 at full sample average for each country
- 2 countries: US & UK
- Time frame: 1990:Q1 2021:Q3

#### Unobserved Component Model

$$100 * \ln \frac{Credit}{GDP} = y_t = \tau_{yt} + c_{yt}$$
(1)

$$100 * InHPI = h_t = \tau_{ht} + c_{ht} \tag{2}$$

• Trends: 
$$\tau_{yt}$$
 &  $\tau_{ht}$ 

$$\begin{aligned} \tau_{yt} = \mu_{yt-1} + \tau_{yt-1} + \eta_{yt}, & \eta_{yt} \sim iidN(0, \sigma_{\eta y}^2) \\ \mu_{yt} = \mu_{yt-1} + \eta_{\mu yt}, & \eta_{\mu yt} \sim iidN(0, 0.01) \\ \tau_{ht} = \mu_{ht-1} + \tau_{ht-1} + \eta_{ht}, & \eta_{ht} \sim iidN(0, \sigma_{\eta h}^2) \\ \mu_{ht} = \mu_{ht-1} + \eta_{\mu ht}, & \eta_{\mu ht} \sim iidN(0, 0.01) \end{aligned}$$

#### Unobserved Component Model

• Cycles: c<sub>yt</sub> & c<sub>ht</sub>

$$c_{yt} = \phi_y^1 c_{yt-1} + \phi_y^2 c_{yt-2} + \phi_y^{\times 1} c_{ht-1} + \phi_y^{\times 2} c_{ht-1} + \varepsilon_{yt}$$
(3)  
$$\varepsilon_{yt} \sim iidN(0, \sigma_{\varepsilon y}^2)$$
(4)

$$c_{ht} = \phi_h^1 c_{ht-1} + \phi_h^2 c_{ht-2} + \phi_h^{\times 1} c_{yt-1} + \phi_h^{\times 2} c_{yt-1} + \varepsilon_{ht}$$
(5)  
$$\varepsilon_{ht} \sim iidN(0, \sigma_{\varepsilon h}^2)$$
(6)

#### Covariance Matrix

(7)

#### Optimization process

• Kalman filter with adjusted Likelihood function:

$$I(\theta) = -0.5 \sum_{t=1}^{T} ln[(2\pi)^2 |f_{t|t-1}|] - 0.5 \sum_{t=1}^{T} \eta'_{t|t-1} f_{t|t-1}^{-1} \eta_{t|t-1} - w1 \sum_{t=1}^{T} (c_{yt}^2) - w2 \sum_{t=1}^{T} (c_{ht}^2)$$

## **Empirical Results**

# VAR(2) - 1 Cross-lag Model Estimate - UK and US

|  |                       | UK VAR2 1-cross lag |                      | US VAR2 1-cross lag |                      |
|--|-----------------------|---------------------|----------------------|---------------------|----------------------|
| Description  | Para.                 | Median              | [10%, 90%]           | Median              | [10%, 90%]           |
| Credit to household 1st AR parameter                       | $\phi_{v}^{1}$        | 1.4238              | [1.3585, 1.4892]     | 1.2074              | [1.1374, 1.2785]     |
| Credit to household 2nd AR parameter                       | $\phi_v^2$            | -0.4698             | [-0.5305, -0.4090]   | -0.2483             | [-0.3152, -0.1825]   |
| Credit to household 1st cross cycle AR parameter           | $\phi_{y}^{\times 1}$ | 0.0238              | [0.0154, 0.0319]     | 0.0318              | [0.0228, 0.0407]     |
| Credit to household 2nd cross cycle AR parameter           | $\phi_y^{\times 2}$   |                     |                      |                     | -                    |
| Housing Price Index 1st AR parameter                       | $\phi_h^1$            | 1.3173              | [1.2647, 1.3701]     | 1.8038              | [1.7700, 1.8363]     |
| Housing Price Index 2nd AR parameter                       | $\phi_h^2$            | -0.3315             | [-0.3885, -0.2746]   | -0.8261             | [-0.8605, -0.7903]   |
| Housing Price Index 1st cross cycle AR parameter           | $\phi_h^{\times 1}$   | -0.0173             | [-0.0464, 0.0062]    | 0.0104              | [0.0007, 0.0204]     |
| Housing Price Index 2nd cross cycle AR parameter           | $\phi_h^{\times 2}$   |                     |                      |                     |                      |
| S.D. of permanent shocks to Credit to household            | $\sigma_{ny}$         | 0.2714              | [0.2150, 0.3155]     | 0.2954              | [0.2312, 0.3414]     |
| S.D. of transitory shocks to Credit to household           | $\sigma_{ey}$         | 0.8021              | [0.7699, 0.8376]     | 0.8631              | [0.8287, 0.9012]     |
| S.D. of permanent shocks to Housing Price Index            | $\sigma_{nh}$         | 0.0789              | [0.0742, 0.0845]     | 0.1390              | [0.1222, 0.1618]     |
| S.D. of transitory shocks to Housing Price Index           | $\sigma_{eh}$         | 1.2242              | [1.1886, 1.2613]     | 0.8988              | [0.8641, 0.9355]     |
| Correlation: Permanent credit to household/Permanent HPI   | $\rho_{nynh}$         | 0.0189              | [-0.3049, 0.3393]    | 0.0082              | [-0.3117, 0.3226]    |
| Correlation: Transitory credit to household/Transitory HPI | $\rho_{eyeh}$         | 0.2536              | [0.1713, 0.3337]     | 0.1537              | [0.0399, 0.2619]     |
| Log-likelihood value                                       | llv                   | 578.6200            | [576.1600, 582.1500] | 204.9400            | [202.4200, 208.4500] |

Note:

UK - US Bayesian method random walk Metropolis-Hasting posterior distribution estimates

# VAR(2) - 1 Cross-lag Model Estimate Summary

- The sum of AR parameters of the cyclical components in all three models is smaller, although close to one
- The standard deviation of the shocks in the cycles σ<sub>ei</sub> is much higher than the standard deviation of the shocks to the trend σ<sub>ni</sub> of both credit and housing prices
- Variations in the housing price cyclical components  $\sigma_{eh}$  of the UK are bigger than in the US
- The correlation of the shocks to the cyclical components among the two variables  $\rho_{eyeh}$  suggests that cyclical variation among housing price and household credit is strongly positively correlated

# Cross-country Comparison of Causal Coefficients

|                | $\phi_{\mathbf{y}}^{\mathbf{x}1}$ HPI on Credit |                   | $\phi_h^{\times 1}$ | $\phi_{h}^{\rm x1}$ Credit on HPI |  |  |
|----------------|---|-------------------|---------------------|-----------------------------------|--|--|
| Country        | Median  | [10%, 90%]        | Median              | [10%, 90%]                        |  |  |
| Australia      | 0.0157  | [-0.0093, 0.0412] | 0.0521              | [0.0014, 0.1060]                  |  |  |
| Belgium        | 0.0279  | [0.0013, 0.0559]  | -0.0656             | [-0.0980, -0.0339]                |  |  |
| Canada         | 0.0191  | [0.0032, 0.0332]  | -0.0152             | [-0.0343, 0.0025]                 |  |  |
| Finland        | 0.0080  | [0.0017, 0.0156]  | 0.0085              | [0.0021, 0.0156]                  |  |  |
| France         | 0.0298  | [0.0185, 0.0411]  | -0.0643             | [-0.1098, -0.0241]                |  |  |
| Germany        | 0.0728  | [0.0500, 0.0917]  | -0.0061             | [-0.0282, 0.0052]                 |  |  |
| Hong Kong      | -0.0031   | [-0.0079, 0.0019] | -0.0629             | [-0.0836, -0.0453]                |  |  |
| Italy          | 0.1001  | [0.0895, 0.1063]  | -0.0027             | [-0.0072, 0.0014]                 |  |  |
| Japan          | -0.0088   | [-0.0326, 0.0174] | 0.1659              | [0.1202, 0.2173]                  |  |  |
| Netherlands    | 0.0058  | [-0.0039, 0.0166] | -0.0043             | [-0.0156, 0.0070]                 |  |  |
| New Zealand    | 0.0078  | [-0.0035, 0.0199] | -0.0139             | [-0.0249, -0.0036]                |  |  |
| Norway         | 0.0109  | [0.0097, 0.0116]  | 0.0059              | [0.0047, 0.0066]                  |  |  |
| South Korea    | 0.0106  | [-0.0033, 0.0308] | 0.0027              | [-0.0251, 0.0369]                 |  |  |
| Spain          | 0.0144  | [0.0003, 0.0331]  | 0.0051              | [-0.0023, 0.0146]                 |  |  |
| Sweden         | 0.0159  | [0.0071, 0.0252]  | 0.0400              | [0.0218, 0.0617]                  |  |  |
| United Kingdom | 0.0238  | [0.0154, 0.0319]  | -0.0173             | [-0.0464, 0.0062]                 |  |  |
| United States  | 0.0318  | [0.0228, 0.0407]  | 0.0104              | [0.0007, 0.0204]                  |  |  |

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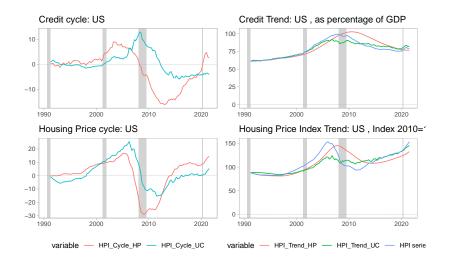
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Cross-country Comparison of Causal Coefficients Summary

- In 11 out of 17 countries, the HPI on Credit causal coefficient  $\phi_y^{\times 1}$  are positive and significant. All 11 countries are in North America and Europe.
- Only 6 countries have positive and significant Credit on HPI causal coefficient  $\phi_y^{\times 1}$ . Three of which have smaller magnitudes than their  $\phi_y^{\times 1}$  counterpart

 $\rightarrow$  Overall, we found evidence that past transitory shocks to house price credit will cause a positive deviation in future transitory household credit. However, the effect in the opposite direction is much smaller and sometimes insignificant

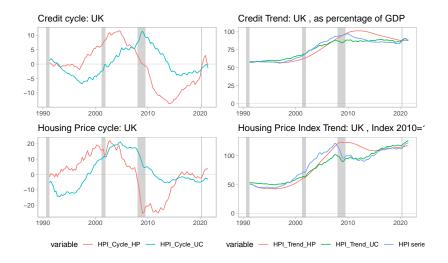
#### Unobserved Component Graphs: United States



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#### Unobserved Component Graphs: United Kingdom



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### Conclusion

- Extracting temporary and permanent components information gave insights on the dynamics of the two series housing and credit in both short-run and long-run
- Evidence showing that past movement of a cycle (HPI) has predictive power over the other cycle (credit)

#### Thank You

I look forward to your questions and comments